

Mutational Robustness and Asymmetric Functional Specialization of Duplicate Genes

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Mutational robustness and asymmetric functional specialization of duplicate genes.

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Abstract: Most duplicate genes are eliminated from a genome shortly after duplication, but those that remain are an important source of biochemical diversity. Much of their diversification arises via functional “specialization”, loss of some functions of the duplicates remaining in the genome. I here present evidence from genome-scale protein-protein interaction data, microarray expression data, and large-scale gene knockout data that this diversification is often asymmetrical: one duplicate usually shows significantly more molecular or genetic interactions than the other. I propose a model that can explain this divergence pattern if duplicate gene pairs are less likely to suffer deleterious mutations when having diverged asymmetrically. The data may provide the first evidence that natural selection has increased mutational robustness in genetic networks.

Soon after a gene duplication, degenerative mutations are likely to eliminate duplicate genes from the genome (Li 1997; Lynch and Conery 2000). However, gene duplications occur continuously and at high rates in eukaryotes, which accounts for the fact that up to 50% of a eukaryotic genome may consist of duplicate genes (Lynch and Conery 2000; Rubin *et al.* 2000). These persisting duplicate genes are perhaps the most prominent source of biochemical innovation of gene products. However, little is known about how this innovation occurs or about how gene duplicates diverge in general.

Studying functional divergence among duplicate genes requires a definition of gene function, but no universal such definition is possible. The reason is that there are several complementary ways of classifying gene functions. For instance, gene products can be characterized biochemically, e.g., as enzymes or transcription factors. Second, they can be characterized through their time and locus of expression, e.g., expression during a cell cycle stage, in the cytoplasm, or during brain development. Third, they can be characterized genetically through mutations and through other genes that these mutations affect. This list is not necessarily complete.

Functional genomics has added much information to each of these categories, especially in model organisms like the yeast *Saccharomyces cerevisiae*. First, monitoring expression through microarrays (Chu *et al.* 1998; Eisen *et al.* 1998; Gasch *et al.* 2000; Spellman *et al.* 1998) provides spatiotemporal expression information for thousands of genes at once. This information is indicative of the biological process a gene is involved in. Second, genome-wide protein-protein interactions can characterize physical interactions among thousands of gene products (Bartel *et al.* 1996; Fromont-Racine *et al.* 1997; Ito *et al.* 2001; Uetz *et al.* 2000). Third, large-scale gene knockout screens in combination with microarray experiments indicate which genes' expression level are affected by a mutated gene. Thus, even in the absence of a detectable phenotype – all too frequent in knock-out experiments – a putative function can sometimes be assigned using genetic interactions with known genes.

Attempts to identify gene functions according to any of the above criteria, whether they use genomic or pre-genomic techniques, yield one key message: most genes have more than one, if not many functions. They are expressed at multiple times and in multiple places, they affect multiple biological processes when mutated, or they interact with proteins with diverse biochemical and biological roles (Bender *et al.* 1983; Gerhart and Kirschner 1998; Jack and Delotto 1995; Kirchhamer *et al.* 1996; Li and Noll 1994; Schwikowski *et al.* 2000; Slusarski *et al.* 1995; Wagner 2001). This multi-functionality has important implications for the divergence of

duplicate genes: duplicate genes often diverge through functional specialization, loss of complementary (sub)functions in each duplicate (Force *et al.* 1999; Lynch and Force 2000; Wagner 2000b). Examples abound. To name but two, the *ZAG1* and *ZMM2* genes are paralogues in the maize genome. They are orthologues of the Arabidopsis *AGAMOUS* gene which is involved in carpel and stamen development. Each of them appears to have largely lost one of their ancestral expression domains: *ZAG1* is expressed at high levels in developing carpels and *ZMM2* is expressed in developing stamens. A null-mutation in *ZAG1* affects only early carpel development. (Coen and Meyerowitz 1991; Mena *et al.* 1996; Schmidt *et al.* 1993). Force and collaborators (Force *et al.* 1999) report on the zebrafish *engrailed* genes *eng1* and *eng1b*, the likely results of a teleost-specific gene duplication of the tetrapod *En1* gene. In mice and chicken, *En1* is expressed in the developing pectoral appendage bud and in specific neurons of the developing hindbrain and spinal cord. In zebrafish, *eng1* retained expression in the pectoral appendage bud, whereas *eng1b* is only expressed in the hindbrain and the spinal cord. Similar patterns of divergence may be quite common in zebrafish (Egger *et al.* 1997; Egger *et al.* 1995; Lee *et al.* 1996).

Studies focussing on individual gene pairs fall short of identifying general divergence patterns of many duplicate genes. At first sight, analyzing functional divergence of many duplicate genes may seem like a hopeless task. Because it is not even straightforward to *classify* one gene's function, how would one *compare* the functions of many divergent duplicates? Functional genomic experiments provide a crude remedy for this problem. Despite their disadvantage of providing largely qualitative information about genetic and molecular interactions of genes, their great advantage is that they do so for thousands of genes at once. They thus yield insight about one aspect – however minute – of gene function, such as the protein interaction partners of a gene, gene expression patterns affected through mutating a gene, or the response of gene expression to environmental challenges. It is this aspect of gene function I will focus on.

Methods

Gene duplication data. Data on yeast gene duplicates was kindly provided by John Conery (University of Oregon, Department of Computer Science) and was generated as described in (Lynch and Conery, 2000). Briefly, gapped BLAST (Altschul *et al.* 1997) was used for pairwise amino acid sequence comparisons of all yeast open reading frames as obtained from GenBank. All protein pairs with a BLAST alignment score greater than 10^{-2} were retained for further analysis. Then, the following conservative approach was followed to retain only unambiguously aligned sequences. Using the protein alignment generated by BLAST as a guide, a sequence pair was scanned to the right of each alignment gap. All sequence from the end of the gap through the first “anchor” pair of matched amino acids was discarded. All subsequent sequence (exclusive the anchor pair of amino acids) was retained if a second pair of matching amino acids was found within less than six amino acids from the first. This procedure was then repeated to the left of each alignment gap (see Lynch and Conery 2000 for more detailed description and justification). The retained portion of each amino acid sequence alignment was then used jointly with DNA sequence information to generate nucleotide sequence alignments of genes. For each gene pair in this data set, the fraction K_s of synonymous (silent) substitutions per silent site, as well as the fraction K_a of replacement substitutions per replacement site were estimated using the method of Li (Li 1993).

Protein interaction data and analysis. Data for 899 pairwise interactions among 985 yeast proteins, as reported in (Uetz *et al.* 2000), was obtained from <http://depts.washington.edu/sfields/projects/YPLM/Nature-plain.html> on February 15, 2000. There are 43 proteins that were reported to interact with themselves. Before further analysis all such self-interactions were eliminated. (Self-interactions are interactions between two protein products of the same gene, such as might occur for homodimerizing proteins.) The resulting protein interaction network was then represented as a graph using LEDA (Mehlhorn and Naher 1999). Within this graph representation, common and different protein interactions among gene family members are easily analyzed (Wagner 2001). To analyze protein interaction data not generated by two-hybrid experiments, I used information on physical interactions among yeast proteins obtained from the MIPS database (Mewes *et al.*, 1999, <http://mips.gsf.de/proj/yeast/CYGD/db/index.html>). I eliminated from this data all protein interactions confirmed by two-hybrid experiments. The remaining 899 interactions involve 680 proteins. I did not distinguish between genes with only one paralogue and genes that occur in multigene in the analysis of either data set.

I tested the observed pattern of the number of interactions for products of paralogous yeast genes against a null-hypothesis of symmetric divergence through *loss-of-interactions*. According to this hypothesis, for two proteins P and P* that (i) are products of duplicate genes, that (ii) have d_1 and d_2 protein interactions, respectively, and that (iii) share b of these interaction partners (Fig. 1a), the d_1-b and d_2-b non-shared interactions of P and P* are due only to loss of interactions after the duplication. In other words, immediately after the gene duplication leading to P and P*, each protein had d_1+d_2-b interactions with the same proteins. P lost d_2-b of these interactions, whereas P* lost d_1-b interactions. Central to this null-hypothesis is that both proteins have equal probability to lose an interaction, that is, that the total number of lost interactions is partitioned between the two proteins according to a binomial distribution $B(d_1+d_2-2b, 1/2)$. The many paralogous gene pairs to be analyzed have different and usually small values for d_1 , d_2 , and b . Thus, it is most expedient to test the null-hypothesis by numerically generating an expected distribution of divergence values (Figs 2b, 3c). This distribution can then be compared with the observed distribution.

A second null-hypothesis is that of symmetric divergence through *gain-of-interactions*. It assumes that immediately after duplication, the two proteins P and P* had only b interactions, and that P gained d_1-b interactions, whereas P* gained d_2-b interactions. Symmetric divergence means that each protein is equally likely to gain an interaction, i.e., that the total number of gained interactions is partitioned between the proteins with a binomial distribution $B(d_1+d_2-2b, 1/2)$.

Environmental stress and gene expression. To assay the differential expression response of yeast paralogues to environmental stresses, I used data provided by Gasch and collaborators (Gasch *et al.*, 2000; http://www-genome-stanford.edu/yeast_stress) for the following conditions: heat shock (25°C to 37°C, after 30 minutes), reverse heat shock (37°C to 25°C, 30'), H₂O₂ and Menadione exposure, both of which generate reactive oxygen species (60' and 80', respectively), dithiothreitol, a reducing agent interfering with protein folding (90'), diamide, an agent oxidizing sulfhydryl groups, (40'), hyperosmotic shock mediated by 1M sorbitol (60'), hypo-osmotic shock mediated by transfer of cells from 1M sorbitol to medium lacking sorbitol (30'), amino acid starvation (2 hours), nitrogen depletion (1d), and stationary phase (7d). I considered genes whose expression level was changed at least three-fold relative in response to a stressor to be affected significantly. Because the expression response to most environmental stresses is transient, I chose a time point (indicated above in parentheses) approximately halfway through the measured response time series for each environmental stress to assess significant change. I then counted the number of stressors to which each member of a paralogous gene pair responded, and did so for all duplicate pairs with $0.5 < K_s < 3$. Due to cross-hybridization, very

closely related duplicates can not be distinguished through microarray analysis. However, the analysis of Gasch (Gasch *et al.* 2000, Fig. 5) suggests that gene pairs with $K_s > 0.5$ are readily distinguishable. I excluded paralogues with $K_s < 0.5$ from the analysis. For 2.6% of all paralogous gene pairs, there was at least one stress condition where the expression of one gene was induced and that of the other was repressed. I excluded all such gene pairs from the analysis.

Gene perturbations and gene expression. Data summarizing the effects of 271 gene deletions (and other treatments) on gene expression was made available as supplemental material to (Hughes *et al.* 2000), file ‘data_expts_1-300_ratios.txt’. From this data set, which contains \log_{10} -transformed expression ratios of 6312 genes for each mutation, I eliminated all data derived from haploid and aneuploid deletion strains, as well as data on non-genetic treatments. The remaining data contains information on null-mutation effects for a total of 21 paralogous gene pairs, the most closely related 11 of which ($K_s < 3.2$) are discussed in here. For each member gene of each paralogue, I determined what other genes were affected in their expression level by a synthetic-null mutation in the gene. I also determined the number of genes that were affected by a null-mutation in each paralogue. I considered a gene as affected by a null-mutation if its level of mRNA expression had changed by more than 3-fold in response to the mutation.

The tests of two null-hypothesis of divergence proceed analogously to those above, except that because of the larger number of genes affected, a meaningful test is possible for each gene pair individually. Let d_1 and d_2 be the number of genes affected by a synthetic null-mutation in gene 1 and 2, respectively, of a paralogous pair. Let b be the number of genes affected by both mutations. The first null-hypothesis is that all divergence is due to equiprobable loss-of-effects on other genes. An exact test is most expedient in this case. It assumes that immediately after the duplication, a null mutation in either duplicate would have affected $d_1 + d_2 - b$ other genes. $l_1 = d_2 - b$ and $l_2 = d_1 - b$ of these effects were subsequently lost in gene 1 and 2, respectively, adding to a total of $d_1 + d_2 - 2b$ lost effects. A disparity between l_1 and l_2 indicates asymmetry in divergence. The probability \mathbf{P} of a disparity as big or bigger as that actually observed, by chance alone, is calculated by summing over the tails of a binomial distribution $B(d_1 + d_2 - 2b, 1/2)$, so

$$\mathbf{P} = 2 \sum_0^{\min(l_1, l_2)} \binom{d_1 + d_2 - 2b}{1/2} \left(\frac{1}{2}\right)^{d_1 + d_2 - b} \quad l_1 \neq l_2,$$

where $\mathbf{P}=1$ for $l_1=l_2$. The factor 2 in front of the summation sign indicates that this is a two-tailed test. The second null-hypothesis, that all divergence is due to equiprobable gain-of-effects by the

two duplicates, is tested in the same way. The only difference is that $\min(l_1, l_2)$ above is replaced with $\min(g_1, g_2)$, where $g_1 = d_1 - b$ and $g_2 = d_2 - b$ is the number of effects gained.

Results and Discussion

Asymmetric divergence in protein-protein interactions. Genome-scale screens of protein interactions using the yeast two-hybrid assay have been carried out in several organisms (Bartel *et al.* 1996; Fromont-Racine *et al.* 1997; Ito *et al.* 2000; Uetz *et al.* 2000). Their results are comprehensive maps of protein-protein interactions comprising most proteins encoded by a genome. Interpreting these maps is still difficult, because they may contain significant numbers of false positive and false negative interactions (Ito *et al.* 2001), and because they collapse the spatial and temporal dimensions of gene expression into a still-life image of protein interactions. However, these maps have also demonstrated usefulness in predicting the spatial expression domain and functional annotation of many proteins from their interaction partners (Schwikowski *et al.* 2000). They can also answer questions about global patterns of interactions, questions whose answer does not depend on the veracity of each individual interaction, but only on statistical interaction patterns.

Gene duplications are important in shaping protein interaction networks. For instance, more than 50% of yeast genes whose products interact with proteins are part of gene families (Wagner 2001). How do gene duplications affect protein interactions? Fig. 1 shows a hypothetical protein P that interacts with four other proteins. Immediately after duplication of the gene encoding P, P and its duplicate P* share all four interactions. As the duplicates diverge in sequence, they also diverge in their protein interactions. Each protein may occasionally gain new interactions. But if mutations are more likely to cause loss of an interaction, as suggested by the prevalence of degenerative mutations in general (Li 1997), then most divergence will be due to loss of originally common protein interactions. Here, I use the number of interaction partners a protein has as a crude one-dimensional indicator of protein function. The number of common and different interactions between two duplicates then indicates their functional divergence.

Fig. 2a shows the number of interaction partners for 2185 pairs of paralogous genes in the network described by Uetz and collaborators (Uetz *et al.* 2000). These comprise all paralogous pairs with $K_s < 3$ synonymous substitutions per synonymous site, corresponding - albeit with a large error margin - to genes duplicated within the last 300Myrs (Wagner 2001). The abscissa and ordinate axes show the number of protein interactions for the first and second protein member of each pair. The number of common interactions in these pairs is small: even among the most

recent paralogues ($K_s < 0.5$) less than 60 percent share any interactions at all, and this number dwindles to less than 15 percent for more distant paralogues ($K_s > 1$) (Wagner 2001).

Fig. 2a shows a distinct L-shape, indicating that in many protein pairs, where one partner has many interactions, the other one has disproportionately few. This negative correlation in the number of interaction partners between duplicates is statistically highly significant (Spearman $r_s = -0.62$, $P < 10^{-3}$; Pearson $r = -0.185$, $P < 10^{-3}$, $df = 2183$). Could it have occurred by chance alone, that is, through random equiprobable loss of interactions in either member of a pair? Fig. 2b shows the results of testing this hypothesis, a hypothesis based on three assumptions. First, all observed differences in interactions among paralogues are due to loss of originally common interactions. Second, each protein interaction of the preduplication state is lost with equal probability in each duplicate (symmetric divergence). Third, the total number of lost interaction per gene pair is the same as that observed empirically. The plot shown in Fig. 2b stems from a stochastic simulation of the divergence of 2185 genes using these assumptions. The L-shape of the plot in Fig. 2a disappears in this scenario of symmetric divergence, as does the highly negative statistical association (Spearman $r_s = -0.1$, $P < 10^{-3}$; Pearson $r = 0.44$, $P < 10^{-3}$, $df = 2183$).

Another possible, although less likely null-hypothesis of divergence, is the independent and equiprobable (symmetric) *gain* of interactions in the two paralogues. Simulating the divergence of paralogues under this null-hypothesis also does not reproduce the characteristic L-shape and the highly negative statistical association of the empirical data (Fig. 2c; Spearman $r_s = -0.1$, $P < 10^{-3}$; Pearson $r = 0.42$, $P < 10^{-3}$, $df = 2183$).

Independent genome-scale two-hybrid experiments using different experimental designs (Ito *et al.* 2001; Uetz *et al.* 2000) show limited overlap in the interactions they detect. It is thus advisable to assure that the observed patterns of divergence are not artefacts of a particular experimental technique. I have repeated the above analysis with yeast protein interaction data taken from the MIPS database (Mewes *et al.* 1999), from which I eliminated all protein interaction information generated by two-hybrid experiments. The remaining 899 interactions among 680 yeast proteins have been experimentally confirmed using techniques ranging from Western blotting to coimmunoprecipitation. The global pattern of interactions among paralogues follows closely that of the two-hybrid data, an L-shaped distribution indicating asymmetry (Fig. 3a) and a highly negative statistical association (Spearman $r_s = -0.57$, $P < 10^{-3}$; Pearson $r = -0.2$, $P < 10^{-3}$, $df = 1803$). This pattern is not explicable through either symmetric loss of interactions (Fig. 3b; Spearman $r_s = 0.11$, $P < 10^{-3}$; Pearson $r = 0.5$, $P < 10^{-3}$, $df = 1803$) or symmetric gain of interactions (Fig. 3c; Spearman $r_s = 0.06$, $P < 10^{-3}$; Pearson $r = 0.48$, $P < 10^{-3}$, $df = 1803$).

In sum, protein interactions among products of duplicate genes diverge asymmetrically, i.e., one paralogue has more protein interactions than the other. This asymmetry is statistically highly significant and is neither explicable through independent (equiprobable) loss of function in the duplicates, nor through independent gain of function.

Asymmetric response to environmental stresses. Unicellular organisms like yeast have evolved elaborate cellular responses allowing them to adapt to drastic environmental changes. They can not only withstand fluctuations in temperature, osmolarity, environmental acidity, and types and quantity of nutrients. They can also survive the influence of radiation and toxic chemicals. During environmental change, many genes alter their transcriptional activity. Such changes in mRNA expression profile provide valuable insights into gene functions (Chu *et al.* 1998; Eisen *et al.* 1998; Gasch *et al.* 2000; Spellman *et al.* 1998). A recent study examined the genomic mRNA expression response of most yeast genes to a variety of environmental stressors (Gasch *et al.* 2000). To assess the differential response of duplicate genes to these stressors, I analyzed data from 11 different stress-responses, including heat shock, hyperosmotic shock, amino acid, and nitrogen starvation (Gasch *et al.* 2000). I excluded the most closely related paralogues ($K_s < 0.5$) from the analysis, because cross-hybridization does not allow them to be distinguished by microarray analysis. I also excluded paralogues where neither member gene responded to any environmental stressor. For the remaining 3815 paralogous gene pair with $0.5 < K_s < 3$, I identified the number of stressors to which each member of the pair responds.

There is again a pronounced asymmetry in the response of gene duplicates to these stresses, as indicated by a significantly negative statistical association between the number of stresses the first and second gene respond to (Spearman $r_s = -0.33$, $P < 10^{-3}$; Pearson $r = -0.12$, $P < 10^{-3}$, $df = 3813$). Completely analogous to the tests for symmetric divergence in protein interactions, I analyzed whether this association is consistent with the null-hypothesis that the paralogues originally responded identically to these 11 stresses, but that some responses got lost and did so with equal probability between the duplicates. This null-hypothesis of symmetric divergence must be rejected (Spearman $r_s = -0.0094$, $P > 0.5$; Pearson $r = 0.18$, $P < 10^{-3}$, $df = 3813$). Nor is the observed pattern explicable if the paralogous genes gained stress responses independently and equiprobably since the duplication (Spearman $r_s = -0.0087$, $P > 0.5$; Pearson $r = 0.18$, $P < 10^{-3}$, $df = 3812$).

In sum, the distinct asymmetry in divergence observed for protein interactions also holds for another aspect of gene function, the response to environmental stress.

Asymmetric response to genetic perturbations. The results of a large-scale gene perturbation experiment in yeast, involving several hundred gene-knockout mutations in

combination with microarray measurements of changes in the expression of 6312 yeast genes, have been reported (Hughes *et al.* 2000). Measuring the effect of a null mutation in a gene on the expression of all other genes does not distinguish between direct and indirect effects of the mutation. Its advantage, however, is that it is a very comprehensive means to assay genetic interactions.

For the purpose of this paper it is relevant that the available data (Hughes *et al.* 2000) contains information on the knockout effect of 11 moderately distant paralogous gene pairs with $K_s < 3.5$. For these 11 gene pairs, I compared the number of genes whose expression is affected by a null-mutation in each member of the pair (Table 1). Interpreting differences between paralogues in the number of affected genes is complicated because these differences are not only the result of divergence between the paralogues. They also include effects from the divergence of genes interacting with each paralogue. However, the advantage of a perturbation approach is that it provides a more comprehensive assessment of functional differences between paralogues than a mere analysis of direct physical protein interactions. It exposes how the effects of a mutation ripple through a transcriptional regulation network.

Analogous to the analysis above, one can ask whether the observed differences between paralogues can be attributed to independent and equiprobable loss of genetic interactions, or to independent gain of interactions since the duplication. For 7 out of 11 gene pairs in Table 1, both these null-hypotheses must be rejected. That is, these seven gene pairs show statistically significant asymmetries in divergence. Eliminating one of two paralogous genes affects a substantially greater number of other genes than eliminating the other.

Asymmetric divergence and mutational robustness. The number of genes whose expression is affected by gene perturbations is moderately large. This makes it possible to derive statistical evidence for asymmetric divergence of paralogous genes from individual gene pairs. Seven out of 11 perturbed gene pairs show such evidence. The number of environmental stresses to which a gene responds is typically smaller. And so is the number of protein interactions of gene products. These smaller numbers make it more difficult to derive solid evidence for asymmetric divergence from individual gene pairs. However, such evidence emerges when analyzing multiple gene pairs.

It can not be excluded that some divergence among duplicates arises through mutations causing gain of interactions, gain of stress responses etc. However, shortly after a duplication, such mutations are probably rare compared to degenerative loss-of-function mutations (Force *et al.* 1999; Li 1997). Consider the example of a gene's transcriptional response to a stressor. Given the small size of yeast transcriptional regulatory regions, mutations are much more likely to

eliminate an existing transcription factor binding site than to generate a new site mediating a stress response. Divergence of functions through gain of transcription factor binding sites—while probably occurring occasionally—must be much less frequent than loss of binding sites. I will thus focus on a simple model of divergence through loss of common functions. Figure 4 explains the basic idea behind this model. It applies to the divergence of genes that have several suitably defined functions (represented by white boxes in Fig. 4a), as indicated by observed molecular interactions or patterns of gene expression. Immediately after a duplication, two duplicates are identical in all these functions. The model makes only two assumptions about the process of divergence, both of them very simple. First, every function must be exercised by at least one of the two genes. Organisms where this does not hold will suffer reduced fitness. Second, a loss-of-function mutation (i) affects each of the duplicates with equal probability ($1/2$), and (ii) eliminates one of the affected gene's functions. In this context, what is the probability P_{Δ} of suffering a deleterious mutation if the two duplicates have diverged symmetrically vs. asymmetrically? Asymmetric divergence means that one duplicate has lost more functions than the other. Assume that since the duplication, duplicates 1 and 2 have lost a fraction l_1 and l_2 of their functions, respectively ($0 < l_1, l_2 \leq 1$). Let $l = l_1 + l_2$ be the total fraction of functions lost ($0 \leq l \leq 1$). If no function is allowed to have been lost in both genes, the probability that a mutational loss of one further function has a deleterious effect is equal to

$$P_{\Delta} = \frac{1}{2} \left(\frac{l_2}{1-l_1} \right) + \frac{1}{2} \left(\frac{l_1}{1-l_2} \right).$$

Upon expressing l_1 and l_2 in terms of the total fraction of functions lost, l , using $x := l_1/l$ ($1-x = l_2/l$) this expression becomes

$$P_{\Delta}(x, l) = \frac{1}{2} \left(\frac{1-x}{(1/l)-x} \right) + \frac{1}{2} \left(\frac{x}{(1/l)-1+x} \right).$$

In this expression, a value of $x=0.5$ indicates symmetric divergence. The pertinent feature of $P_{\Delta}(x, l)$ is that it is unimodal: Regardless of l , it has a maximum at $x=0.5$ (Fig. 4b). This means that the probability of a deleterious mutation is greatest if two genes have diverged symmetrically. Thus, asymmetric divergence minimizes the risk of deleterious mutations.

If this model is correct, we see asymmetrically diverged gene pairs because organisms harboring them have survived preferentially in the past. Importantly, natural selection would act in an indirect, “second order” manner on such gene pairs. In a population polymorphic for gene duplicates at different stages of divergence, different individuals would not necessarily have different fitness. Rather, the propensity of such individuals to suffer deleterious mutations would be different. Individuals with symmetrically diverged duplicates would be preferentially eliminated from the population through deleterious mutations.

One might assume that the selective advantage of having asymmetrically diverged gene duplicates must be minute. After all, differences in fitness do not manifest themselves until new loss-of-function mutation arise. For any organism, the expected waiting time for such a new loss-of-function mutation is proportional to the inverse of the mutation rate μ (Hartl and Clark 1988). During this time, symmetrically diverged gene duplicates are free to go to fixation via random drift. Formal population genetic analysis (Wagner 2000b) shows that for sufficiently large population sizes ($N > 1/\mu$), the lens of natural selection has sufficient resolving power to perceive differences in mutational robustness and act on them. For microorganisms like yeast, attainable population sizes may well be in the required range. In addition, this minimally required population size is based on the evolution of only one diverging gene pair (Wagner 2000b). It may be much smaller for multiple gene pairs, and their cumulative effects on mutational robustness.

The requirement for large effective population sizes suggests a test for the model. In organisms with small effective population sizes, such as many higher vertebrates, we would not expect asymmetric divergence of gene duplicates. (The necessary data is not yet available.) A requirement for persistently large population sizes may also be one of the reasons why the asymmetry observed is not perfect and does not hold for all genes. Depending on a gene and its functions, a loss-of-function mutation may have very subtle fitness effects. In conjunction with fluctuating effective population sizes, the selection pressures for asymmetric divergence may fluctuate as well. Some genes thus diverge symmetrically whereas others do not.

Mutational robustness – the insensitivity of a biological process to mutations – is a prominent feature of genetic networks. It occurs in all organisms, from single-celled to mammalian. It occurs in processes as different as metabolism and development. And it has been detected using every approach from classical genetics to functional genomics (Hartl *et al.* 1985; Hartman *et al.* 2001; Rendel 1979; Rutherford and Lindquist 1998; Smith *et al.* 1996; vonDassow *et al.* 2000).

There are two possible evolutionary origins for observed robustness. First, whenever gene products interact to perform a biological function, the resulting gene network and its

function might show some resilience to mutations. Second, mutational robustness may be a result of stabilizing natural selection, sorting more robust genetic networks from less robust ones on enormous time scales. The second hypothesis, if true, has profound consequences on our understanding of biological evolution. It implies that natural selection itself can change the amount of phenotypic variation produced by a genetic network through mutations. It could thus also affect any further response of a population to natural selection – the population's evolvability. Despite these intriguing consequences, empirical evidence for evolved robustness is hard to come by. First, robustness is most likely a global property of genetic networks (Edwards and Palsson 2000; Wagner 2000a), and thus not easily subject to experimental manipulation. Second, if evolution of robustness occurs, it probably occurs on very large time scales. The third reason has to do with the detection of robustness itself. It is impossible to prove by direct experimentation that phenotypic robustness to mutations translates into robust fitness. The reasons are twofold. First, experimentally undetectably small fitness differences in organisms may be perceived by natural selection. Second, it is usually impossible to reconstitute in the laboratory an environment representing all environments in which an organism may need to survive and reproduce.

For all these reasons, our best chance of assessing whether genuine robustness exists and evolves is to detect patterns of molecular evolution consistent with it (Hirsh and Fraser 2001). Patterns like the one detected here may thus be ideally suited to answer two difficult but important questions in one sweep. First, do differences in mutational robustness exist? And second, are they relevant for evolution, are they seen by natural selection? The data presented here suggest the answer to both questions is yes.

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Figure Captions

Fig. 1.: Asymmetric divergence in protein-protein interactions. (a) Circles stand for proteins, lines for interactions among proteins. Shortly after a gene duplication, the products P and P* of a duplicate gene that are part of a protein interaction network interact with the same proteins. Eventually, some or all of the common interactions may be lost, and new interactions may be gained by either protein.

Fig. 2.: Asymmetric divergence in protein-protein interactions: Two-hybrid data (a)

The number of interaction partners of protein 1 versus protein 2, plotted for the two protein products of 2185 paralogous yeast gene pairs with $K_s < 3$. Protein interaction data is taken from (Uetz *et al.* 2000). Values of zero on either axis indicate that one member of a paralogous pair is not part of the protein interaction network. It may have lost all protein interactions (but may have retained biological functions not mediated through protein interactions.) One protein pair with 0 and 24 interactions is outside the scale shown here but was included in the statistical analysis. (b) and (c) The expected distribution of interactions for the same 2185 paralogues if divergence after duplication had occurred through independent equiprobable interaction loss or gain, respectively, in either duplicate.

Fig. 3.: Asymmetric divergence in protein-protein interactions: Non two-hybrid data

(a) The number of interaction partners of protein 1 versus protein 2, plotted for the two protein products of 1805 paralogous yeast gene pairs with $K_s < 3$. Interaction data supported by experiments not using the two-hybrid assay was obtained from the MIPS data base (Mewes *et al.* 1999). Values of zero on either axis indicate that one member of a paralogous pair is not part of the protein interaction network. It may have lost all protein interactions (but may have retained biological functions not mediated through protein interactions.) (b) and (c) The expected distribution of interactions for the same 1805 paralogues if divergence after duplication had occurred through independent equiprobable interaction loss or gain, respectively, in either duplicate.

Fig. 4.: Asymmetric divergence and mutational robustness. (a) Schematic depiction of symmetric vs. asymmetric divergence of two duplicate genes with 20 (sub)functions, represented by white boxes. Black boxes indicate that a gene has suffered a mutational loss of the respective

function. In asymmetric divergence, this loss occurs preferentially in one gene. **(b)** The probability P_{Δ} that a loss of function mutation has a deleterious effect, that is, that it eliminates a function not “covered” by the other gene, as a function of x , the degree of asymmetry in divergence. P_{Δ} is smallest for maximal asymmetry in divergence, i.e., for $x=0$ and $x=1$.

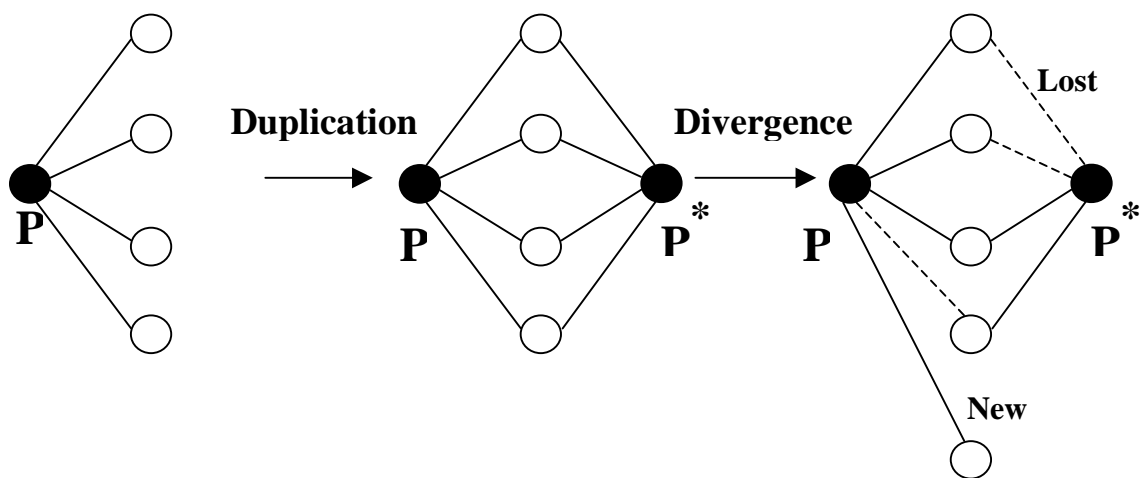


Fig. 1

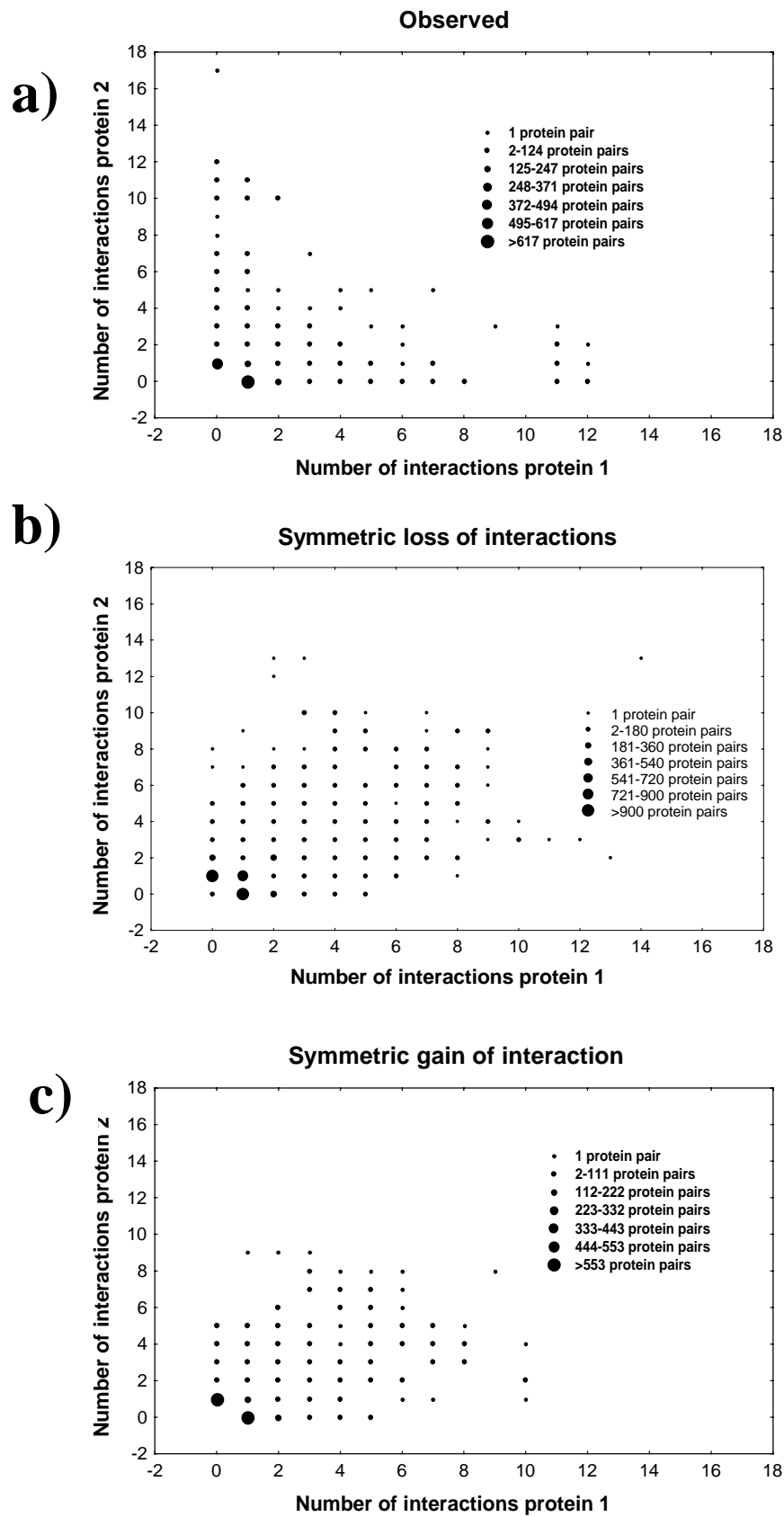
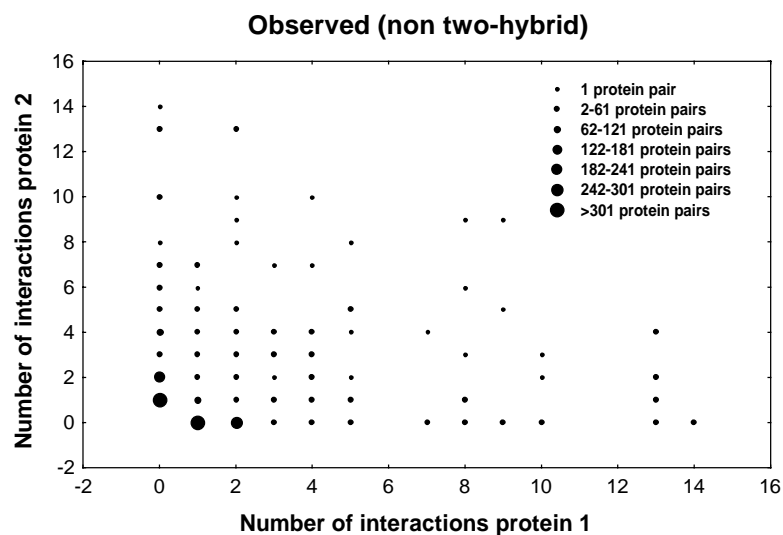
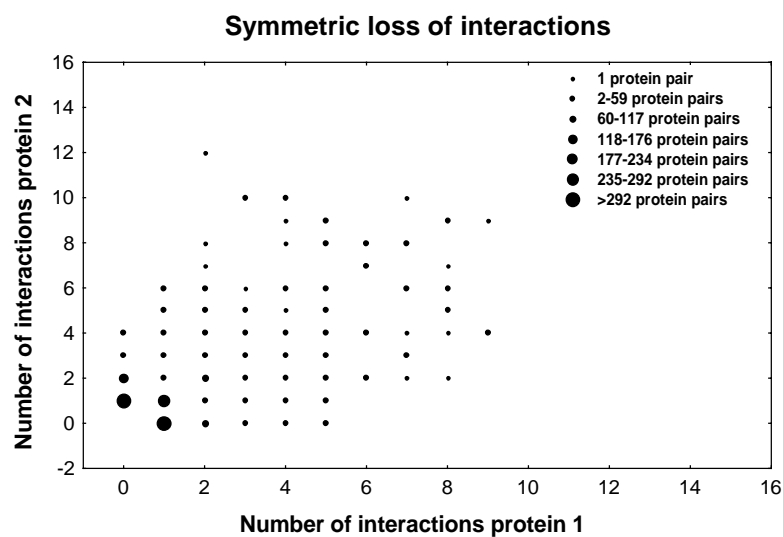


Fig. 2

a)



b)



c)

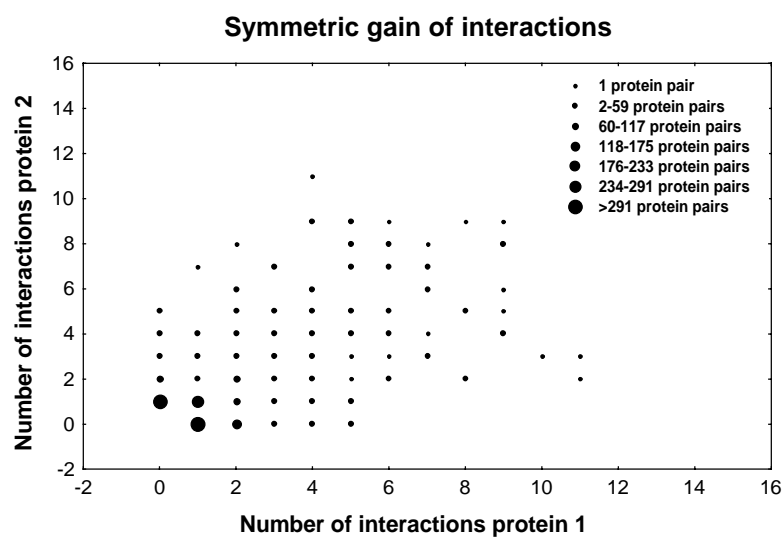
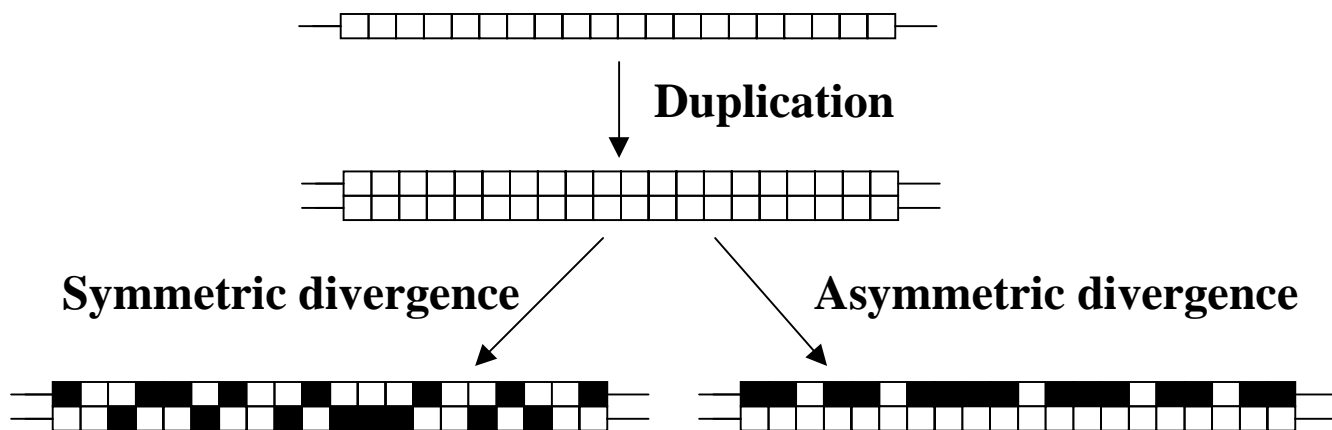


Fig. 3

a)



b)

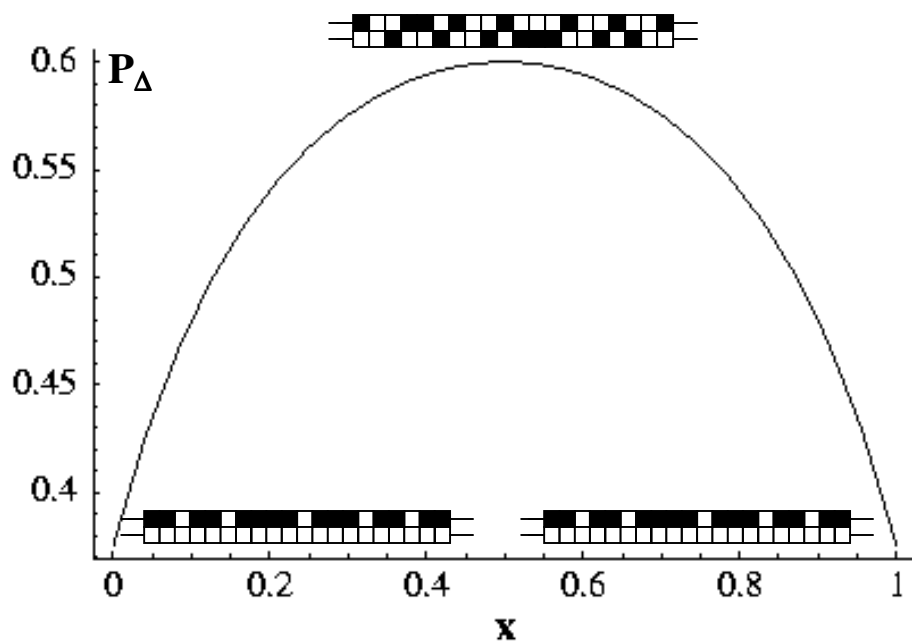


Fig. 4

Table 1: Differential effects of paralogous gene null-mutations on yeast gene expression.

The column labeled '1/2' indicates the number of genes whose expression is affected by a null-mutation in gene 1 and 2, respectively. The column labeled 'common' indicates the number of genes affected by both. **P**-values indicate the probability that a difference in the number of affected genes equal or greater than the observed difference is due to equiprobable loss or gain of effects in the duplicates. Rows in bold type have **P**<0.05.

Gene 1 / Gene 2	K_s	Number of effects		Symmetric Divergence
		1 / 2	Common	
YDL056W / YER111C	1.29	5 / 149	1	7.69×10^{-39}
YAL007C / YOR016C	1.58	6 / 1	0	0.13
YGR109C / YPR119W	1.83	17 / 17	0	1
YBR245C / YOR304W	1.90	20 / 13	3	0.1
YER041W / YKL113C	2.01	0 / 14	0	1.2×10^{-4}
YHR022C / YOR089C	2.08	12 / 3	0	0.035
YLR014C / YMR280C	2.17	14 / 0	0	1.2×10^{-4}
YDL042C / YOR025W	2.36	16 / 97	2	1.9×10^{-16}
YEL049W / YOR009W	2.36	1 / 10	0	0.011
YER073W / YHR039C	2.64	6 / 4	0	0.75
YDR480W / YPL049C	3.21	9 / 23	2	5.2×10^{-3}